# **BE.104 Spring Biostatistics: Detecting Differences and Correlations J. L. Sherley**

### Outline

- 1) Review Concepts
- 2) Detecting differences and quantifying confidence
- 3) Detecting relationships and quantifying confidence

<u>Variance =  $\sigma^2$ </u>

What do changes in variance tell us? (Review in-class exercises)

<graphs>

Multiple "populations" present Skewed data; non-normal data

An important distinction about the application of normal statistics that is often confused:

The sampled **POPULATION** should be normally distributed- why?

Question: If a sample distribution is not normal can we apply parametric statistical methods?

Yes, if they are a sample from an "ideal population" that <u>is</u> normally distributed. It is the properties of the ideal population that <u>matter</u>, not the distribution of the sample, per se. Caveat?

<graph>

# Parametric statistical methods address the uncertainty of sampling.

Now we focus on the structure of the sample because we know it gives us <u>some</u> information about the structure of the ideal population.

So, we must base our decision about using normal statistics on the <u>sample</u> when we have no <u>a priori</u> information about the structure of the ideal population with N members.

### Now to detecting differences

What we want to ask is:

Are two means more different than we would expect based on "error" & statistical variation alone?

Consider there is only one ideal population, and we may be looking at sampling variation or statistical variation and error.

<picture>

So, we define the possible range of differences in the means that could occur by chance/error with some level of confidence.

<graph>

If this difference is not explained by error & statistical variation (i.e., variance) we then can consider other factors:

E.g., A change in a physiological mechanism

Which might lead us to: Parallel testing (i.e., a bigger study) Orthogonal testing (i.e. different kind of study; intervention experiment) "More on this later"

"Detecting Differences Between Estimated Pop. Means"

Our question can be phrased this way regarding the ideal population thinking:

<picture with graph>

Two sample distributions, with numerical different sample means, drawn from ideal populations A and B. Hypothesis: A and B are distinct populations with distinct population means,  $\mu$ .

The null hypothesis: the observed numerical difference occurs due to variance, and the two sample distributions actually derive from the same population.

What do we need to consider for this evaluation?

Given:

1) the magnitude of x  $bar_A - x bar_B$ 2) the variance about x  $bar_A + x bar_B$ 

3) Variance =  $\sigma^2 \approx s^2$ 

Different ways to state the question:

What is the probability that you will call them different when they are not?

What is the probability of being wrong when you think that Pop<sub>A</sub> and Pop<sub>B</sub> are different?

What is the probability that the observed numerical difference in sample means occurs due to chance & errors when population means are equal ( $\mu_A = \mu_B$ ); i.e., there is really only one population that the sample is drawn from?

William Sealy Gosset developed a statistic that gives this probability. Published in 1908 in *Biometrika* under the pseudonym "Student"

Student's t-statistic (originally called "z")

 $t = \underbrace{xbar_{A} - xbar_{B}}_{\sqrt{(s_{A}^{2}/n_{A}) + (s_{B}^{2}/n_{B})}} = \underbrace{difference \text{ in sample means}}_{\text{standard error of the}}$ 

t defines the probability that the observed numerical  $xbar_a - xbar_b$  occurs by chance t defines the probability that they are equal when you think that they differ t defines the probability of being wrong when you think the null hypothesis is not supported

The distribution defined by t is a normal distribution

<Develop graphical explanation>

#### t - distribution properties

1) At a given *t*, two ways to be wrong  $\Rightarrow$  "2-tailed test"  $\mu_A$  not >  $\mu_B$ ; and  $\mu_A$  not <  $\mu_B$ 

One tail:  $\mu_A$  not  $\geq \mu_B$  only; or  $\mu_A$  not  $\leq \mu_B$  only

E.g., measurements that cannot be negative versus zero:  $\mu_0 \le \mu_B$  only BE WARY- the one tail statistic gives the same level of significance at 1/2 the power (more later)

- 2) As *t* increases, the probability of being wrong, when you think  $\mu_A \neq \mu_B$ , decreases I.e., large *t* values are GOOD for the hypothesis that there is a difference in the means.
- As n increases, t increases
   Tables allow us to assign <u>confidence levels</u> (p values) for t @ a given n

See *t*- table; Schork and Remington A-5

p is inversely related to t (p = 1- level of confidence that the null hypothesis is not supported)

I.e., small p is GOOD for the hypothesis that there is a difference in the means.

p < 0.05

Predicts that if you performed this comparison (x bar<sub>A</sub> vs x bar<sub>B</sub>) 20

times, you would be wrong about their being different  $\leq 1$  time, when you thought they were from different populations.

"< a 5% probability that an observed numerical difference occurs due to chance/error"

What is the specific method?

- 1) Compute *t* from x bar<sub>A</sub>,  $s_A$ ,  $n_A$  and x bar<sub>B</sub>,  $s_B$ ,  $n_B$
- 2) Go to t-statistic table: df =  $n_A + n_B 2$ , indicator of sample size
- 3) Extract p, probability of being wrong, when you conclude that two samples came from different populations.

#### Other ways to think about the t- stat

## <Graph>

1) Suppose we computed the 95% CI for  $\mu_A$  and  $\mu_B$  from x bar<sub>A</sub>, s<sub>A</sub> and x bar<sub>B</sub>, s<sub>B</sub>?

What can we say if the intervals don't overlap?

What can we say if the intervals do overlap?

2) x bar<sub>A</sub>- x bar<sub>B</sub> is an estimate of the "Population of  $\mu_A$  -  $\mu_B$  values."

Therefore, we can compute 95% confident interval for  $\mu_A$  -  $\mu_B$  about x bar<sub>A</sub>- x bar<sub>B</sub>

95% CI for  $\mu_{\rm A}$  -  $\mu_{\rm B}$  = (x bar<sub>A</sub>- x bar<sub>B</sub>) ± ( $t_{0.05}$ ) ( $\sqrt{(S^2_{\rm A}/n_{\rm A})}$  + ( $S^2_{\rm B}/n_{\rm B}$ ))

What would you conclude if the 95% CI for  $\mu_A - \mu_B$  about x bar<sub>A</sub>- x bar<sub>B</sub> = 2.5 was -0.5 to 5.5 ?

Sometimes we are comparing a data set to a single value. E.g., How does a set of measurements compare to a standard value, sv (not SD and variance!)

<drawing>

$$t = \frac{\text{sv} - \text{x bar}}{\text{s}/\sqrt{n}}$$
95% CI for  $\mu$ -x = (x bar - sv) ±  $t_{0.05}$ (s/ $\sqrt{n}$ )

Similar: determine 95% CI for µ If sv outside of interval? "Outlier Evaluation" If sv in interval?

Paired t-test (versus "unpaired")

Sometimes two sets of data are paired (e.g., heart rate before & after Rx) Advantage- Avoid intra-set variation due to <u>known</u> causes or effects e.g. day-to-day variability (before vs after treatment); lot to lot variation

<drawing> Calculate differences d, d bar, and s<sub>d</sub> from the distributions:  $t_d = (d \text{ bar})/(s_d/\sqrt{n})$ 

Caution- Must be paired <u>a priori</u> to avoid bias! (same time, conditions, etc.) Use when  $s_1$  and  $s_2$  are large. When  $s_1$  and  $s_2$  are already small, pairing reduces the sensitivity of the test. Why?

## Detecting & Quantifying Associations

Encounter lots of potential causative factors, where

Factor  $X \rightarrow Disease/Toxicity$ 

How do we begin to move from conjecture and anecdotes to causes and effect?

Quantify Associations-	Place a quantitative value on <u>how much</u> Factor X is associated with changes in disease/toxicity
1) Concordance analysis-	"X $\rightarrow$ disease/toxicity?"; and "Not X $\rightarrow$ Not disease/toxicity?"
More later on this.	

2) OR, RR ideas developed earlier

3) <u>Correlation</u> -	How does disease/toxicity vary with changes in X?
	1) similar to "dose response"
	2) location, time, persons, factors, too!

<graph>

General Discussion for Detecting Associations or Correlations

Begin with consideration of a population of related variables X and Y, with Y dependent on X suspected. (e.g. diet vs weight).

< X vs Y scatter plot>

We call Y, the dependent variable; and X, the independent variable

The typical problem we confront:

<diagram>

Question: What is the nature of the population from which the sample was drawn?

1) Was there an association between X and Y?

<u>OR</u>

2) Does the <u>observed</u> association occur by error/chance in the sampling?

First Approach

Determine best fit line to the data by linear regression. "Least squares fit"

<Graph>

 $\Sigma (Y_i - \text{line } Y - \text{value})^2$ 

Minimize  $\Rightarrow$  regression line:  $y^{A} = a + bx$ 

b, the slope, is a measure of the degree to which Y depends on X

When b = 0, Y does not depend on X b > 0, Y is positively correlated with X b < 0, Y is negatively correlated with X

Can detect associations with  $1^{st}$  approach, but no quantification of confidence that the data are not so arranged by chance. Our question is <u>not</u>, "How well does the regression line fit the data?" (as e.g., the root mean squared deviation would be).

## It's, "How likely is it that the observed association occurred by chance/error?"

Second Approach

The linear regression approach <u>assumes</u> that dependency is known. I.e., in our example Y depends on X. However, the regression result can be different if the dependency is reversed:

Y vs y<sup> $\wedge$ </sup> variability  $\neq$  X vs x<sup> $\wedge$ </sup> in some cases

Pearson Product- Moment Correlation Coefficient

- 1) Asks: Is there an association when dependency relationship is not established?
- 2) Gives a quantitative measure of the <u>strength</u> of a detected association
- 3) Gives a measure of the confidence that a detected association occurs for reasons other than chance/error

 $r = \underbrace{\Sigma(X-Xbar) (Y-Ybar)}_{\sqrt{\Sigma}(X-Xbar)^2 \Sigma(Y-Ybar)^2}$ 

-1 ≤ r≤1

When r = 0, no association

r = 1, strong **positive** correlation

i.e., 100% of the variation in one variable is accounted for variation in the other r = -1, strong negative correlation

How does r relate to the regression analysis?

$$r = \sqrt{1 - \frac{\sum [Y - (a + bx)]^2}{\sum (Y - Ybar)^2}}$$

$$\frac{V_{residual}}{V_{tot}}, \text{ variance about the regression line}$$

$$V_{tot}, \text{ variance in Y data}$$

$$r = \sqrt{1 - \frac{V_{res}}{V_{tot}}}$$

1) As  $V_{res} \rightarrow 0$ , there is no variation about the regression line  $r \rightarrow 1.0 \Rightarrow Y$  predicted by X w/o uncertainty

2) When  $V_{res} \rightarrow V_{tot}$ ,  $r \rightarrow 0 \Rightarrow$  no trend in data: X cannot be used to predict Y

$$r^2 = 1 - \frac{V_{res}}{V_{tot}}$$
 "coefficient of determination"  
"R<sup>2</sup>" sometimes

 $0 \leq r^2 \leq 1$ 

"Gives an intuitive feel for how well a straight line describes relationship between X and Y." The degree to which Y depends on X.

Statistics have been developed for a given r or  $r^2$  and  $n_x + n_y - 2$  degrees of freedom to determine p value for confidence that the <u>association between x + y does not</u> occur by chance/error. [<u>Not</u>: "How well does the line fit the data?"]