Quiz I Review Probabilistic Systems Analysis 6.041SC

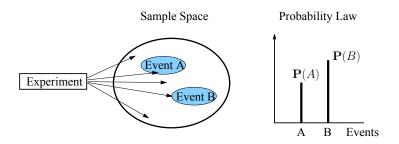
Massachusetts Institute of Technology

Quiz Information

• Content: Chapters 1-2, Lecture 1-7, Recitations 1-7, Psets 1-4, Tutorials 1-3

A Probabilistic Model:

- **Sample Space**: The set of all possible outcomes of an experiment.
- **Probability Law**: An assignment of a nonnegative number **P**(E) to each event E.



Probability Axioms

Given a sample space Ω :

- 1. Nonnegativity: $P(A) \ge 0$ for each event A
- 2. Additivity: If A and B are disjoint events, then

$$\mathbf{P}(A\cup B)=P(A)+P(B)$$

If A_1, A_2, \ldots , is a sequence of disjoint events,

$$\mathbf{P}(A_1 \cup A_2 \cup \cdots) = P(A_1) + P(A_2) + \cdots$$

3. Normalization $P(\Omega) = 1$

Properties of Probability Laws

Given events A, B and C:

- 1. If $A \subset B$, then $\mathbf{P}(A) \leq \mathbf{P}(B)$
- 2. $\mathbf{P}(A \cup B) = \mathbf{P}(A) + \mathbf{P}(B) \mathbf{P}(A \cap B)$
- 3. $\mathbf{P}(A \cup B) \leq \mathbf{P}(A) + \mathbf{P}(B)$
- 4. $\mathbf{P}(A \cup B \cup C) = \mathbf{P}(A) + \mathbf{P}(A^c \cap B) + \mathbf{P}(A^c \cap B^c \cap C)$

Discrete Models

Discrete Probability Law: If Ω is finite, then each event A ⊆ Ω can be expressed as

$$A = \{s_1, s_2, \ldots, s_n\} \qquad s_i \in \Omega$$

Therefore the probability of the event A is given as

$$\mathbf{P}(A) = \mathbf{P}(s_1) + \mathbf{P}(s_2) + \cdots + \mathbf{P}(s_n)$$

• Discrete Uniform Probability Law: If all outcomes are equally likely,

$$\mathsf{P}(A) = rac{|A|}{|\Omega|}$$

Conditional Probability

Given an event B with P(B) > 0, the conditional probability of an event A ⊆ Ω is given as

$$\mathbf{P}(A|B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)}$$

- P(A|B) is a valid probability law on Ω, satisfying

 P(A|B) ≥ 0
 P(Ω|B) = 1
 P(A₁ ∪ A₂ ∪ · · · |B) = P(A₁|B) + P(A₂|B) + · · · ,
 where {A_i}_i is a set of disjoint events
- **P**(*A*|*B*) can also be viewed as a probability law on the restricted universe *B*.

Multiplication Rule

 A_1

• Let A_1, \ldots, A_n be a set of events such that

$$\mathbf{P}\left(\bigcap_{i=1}^{n-1}A_i\right)>0.$$

Then the joint probability of all events is

$$\mathbf{P}\begin{pmatrix} \bigcap_{i=1}^{n} A_i \end{pmatrix} = \mathbf{P}(A_1)\mathbf{P}(A_2|A_1)\mathbf{P}(A_3|A_1 \cap A_2) \cdots \mathbf{P}(A_n| \bigcap_{i=1}^{n-1} A_i)$$

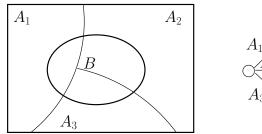
$$A_2 \qquad A_3 \qquad A_1 \cap A_2 \cap A_3$$

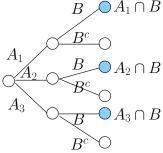
$$\mathbf{P}(A_2|A_1) \mathbf{P}(A_3|A_1 \cap A_2) \qquad A_1 \cap A_2 \cap A_3$$

Total Probability Theorem

Let A_1, \ldots, A_n be disjoint events that partition Ω . If $\mathbf{P}(A_i) > 0$ for each *i*, then for any event B,

$$\mathbf{P}(B) = \sum_{i=1}^{n} \mathbf{P}(B \cap A_i) = \sum_{i=1}^{n} \mathbf{P}(B|A_i)\mathbf{P}(A_i)$$

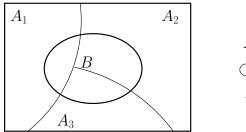


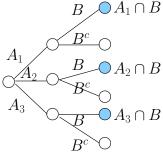


Bayes Rule

Given a finite partition A_1, \ldots, A_n of Ω with $\mathbf{P}(A_i) > 0$, then for each event B with $\mathbf{P}(B) > 0$

$$\mathbf{P}(A_i|B) = \frac{\mathbf{P}(B|A_i)\mathbf{P}(A_i)}{\mathbf{P}(B)} = \frac{\mathbf{P}(B|A_i)\mathbf{P}(A_i)}{\sum_{i=1}^{n}\mathbf{P}(B|A_i)\mathbf{P}(A_i)}$$





Independence of Events

Events A and B are independent if and only if
 P(A ∩ B) = P(A)P(B)

or

$$\mathbf{P}(A|B) = \mathbf{P}(A)$$
 if $\mathbf{P}(B) > 0$

• Events A and B are **conditionally independent** given an event C if and only if

$$\mathbf{P}(A \cap B|C) = \mathbf{P}(A|C)\mathbf{P}(B|C)$$

or

$$\mathbf{P}(A|B \cap C) = \mathbf{P}(A|C)$$
 if $\mathbf{P}(B \cap C) > 0$

• Independence \Leftrightarrow Conditional Independence.

Independence of a Set of Events

 The events A₁,..., A_n are pairwise independent if for each i ≠ j

$$\mathbf{P}(A_i \cap A_j) = \mathbf{P}(A_i)\mathbf{P}(A_j)$$

• The events A_1, \ldots, A_n are **independent** if

$$\mathbf{P}\left(igcap_{i\in S} A_i
ight) = \prod_{i\in S} \mathbf{P}(A_i) \quad \forall \ S\subseteq \{1,2,\ldots,n\}$$

 Pairwise independence ⇒ independence, but independence ⇒ pairwise independence.

Counting Techniques

• **Basic Counting Principle:** For an *m*-stage process with *n_i* choices at stage i,

Choices = $n_1 n_2 \cdots n_m$

• **Permutations:** *k*-length sequences drawn from *n* distinct items without replacement (order is important):

Sequences =
$$n(n-1)\cdots(n-k+1) = \frac{n!}{(n-k)!}$$

• **Combinations:** Sets with *k* elements drawn from *n* distinct items (order within sets is not important):

Sets =
$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

Counting Techniques-contd

• **Partitions**: The number of ways to partition an *n*-element set into r disjoint subsets, with n_k elements in the k^{th} subset:

$$\binom{n}{n_1, n_2, \dots, n_r} = \binom{n}{n_1} \binom{n-n_1}{n_2} \cdots \binom{n-n_1-\dots-n_r-1}{n_r}$$
$$= \frac{n!}{n_1! n_2!, \dots, n_r!}$$

where

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$
$$\sum_{i=1}^{r} n_i = n$$

Quiz I Review

Discrete Random Variables

• A **random variable** is a real-valued function defined on the sample space:

$$X:\Omega \to R$$

• The notation $\{X = x\}$ denotes an event:

$$\{X = x\} = \{\omega \in \Omega | X(\omega) = x\} \subseteq \Omega$$

The probability mass function (PMF) for the random variable X assigns a probability to each event {X = x}:

$$p_X(x) = \mathbf{P}(\{X = x\}) = \mathbf{P}(\{\omega \in \Omega | X(\omega) = x\})$$

PMF Properties

- Let X be a random variable and S a countable subset of the real line
- The axioms of probability hold:

1.
$$p_X(x) \ge 0$$

2. $\mathbf{P}(X \in S) = \sum_{x \in S} p_X(x)$
3. $\sum_x p_X(x) = 1$

• If g is a real-valued function, then Y = g(X) is a random variable:

$$\omega \xrightarrow{X} X(\omega) \xrightarrow{g} g(X(\omega)) = Y(\omega)$$

with PMF

$$p_Y(y) = \sum_{x \mid g(x) = y} P_X(x)$$

Quiz I Review

Expectation

Given a random variable X with PMF $p_X(x)$:

•
$$\mathbf{E}[X] = \sum_{x} x p_X(x)$$

• Given a derived random variable Y = g(X):

$$\mathbf{E}[g(X)] = \sum_{x} g(x)p_X(x) = \sum_{y} yp_Y(y) = E[Y]$$
$$\mathbf{E}[X^n] = \sum_{x} x^n p_X(x)$$

• Linearity of Expectation: $\mathbf{E}[aX + b] = a\mathbf{E}[X] + b$.

Variance

The expected value of a derived random variable g(X) is

$$\mathsf{E}[g(X)] = \sum_{x} g(x) p_X(x)$$

The variance of X is calculated as

• $var(X) = \mathbf{E}[(X - \mathbf{E}[X])^2] = \sum_x (x - \mathbf{E}[X])^2 p_X(x)$ • $var(X) = \mathbf{E}[X^2] - \mathbf{E}[X]^2$ • $var(aX + b) = a^2 var(X)$

Note that $var(x) \ge 0$

Multiple Random Variables

Let X and Y denote random variables defined on a sample space Ω .

• The **joint PMF** of X and Y is denoted by

$$p_{X,Y}(x,y) = \mathbf{P}\big(\{X = x\} \cap \{Y = y\}\big)$$

• The marginal PMFs of X and Y are given respectively as

$$p_X(x) = \sum_{y} p_{X,Y}(x,y)$$
$$p_Y(y) = \sum_{x} p_{X,Y}(x,y)$$

Functions of Multiple Random Variables Let Z = g(X, Y) be a function of two random variables • PMF:

$$p_Z(z) = \sum_{(x,y)|g(x,y)=z} p_{X,Y}(x,y)$$

• Expectation:

$$\mathsf{E}[Z] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$$

• Linearity: Suppose g(X, Y) = aX + bY + c.

$$\mathsf{E}[g(X,Y)] = a\mathsf{E}[X] + b\mathsf{E}[Y] + c$$

Conditioned Random Variables

 Conditioning X on an event A with P(A) > 0 results in the PMF:

$$p_{X|A}(x) = \mathbf{P}\big(\{X = x\}|A\big) = \frac{\mathbf{P}\big(\{X = x\} \cap A\big)}{\mathbf{P}(A)}$$

 Conditioning X on the event Y = y with P_Y(y) > 0 results in the PMF:

$$p_{X|Y}(x|y) = \frac{\mathsf{P}(\{X = x\} \cap \{Y = y\})}{\mathsf{P}(\{Y = y\})} = \frac{p_{X,Y}(x,y)}{p_Y(y)}$$

Conditioned RV - contd

Multiplication Rule:

$$p_{X,Y}(x,y) = p_{X|Y}(x|y)p_Y(y)$$

• Total Probability Theorem:

$$p_X(x) = \sum_{i=1}^n \mathbf{P}(A_i) p_{X|A_i}(x)$$
 $p_X(x) = \sum_y p_{X|Y}(x|y) p_Y(y)$

Conditional Expectation

Let X and Y be random variables on a sample space Ω.
Given an event A with P(A) > 0

$$\mathsf{E}[X|A] = \sum_{x} x p_{X|A}(x)$$

• If $P_Y(y) > 0$, then

$$\mathsf{E}[X|\{Y=y\}] = \sum_{x} x p_{X|Y}(x|y)$$

Total Expectation Theorem: Let A₁,..., A_n be a partition of Ω. If P(A_i) > 0 ∀i, then

$$\mathsf{E}[X] = \sum_{i=1}^{n} \mathsf{P}(A_i) \mathsf{E}[X|A_i]$$

Quiz I Review

Independence

Let X and Y be random variables defined on Ω and let A be an event with $\mathbf{P}(A) > 0$.

• X is independent of A if either of the following hold:

$$p_{X|A}(x) = p_X(x) \ orall x \ p_{X,A}(x) = p_X(x) \mathbf{P}(A) \ orall x$$

• X and Y are independent if either of the following hold:

$$p_{X|Y}(x|y) = p_X(x) \ \forall x \forall y \ p_{X,Y}(x,y) = p_X(x) p_Y(y) \ \forall x \forall y$$

Independence

If X and Y are independent, then the following hold:

- If g and h are real-valued functions, then g(X) and h(Y) are independent.
- $\mathbf{E}[XY] = \mathbf{E}[X]\mathbf{E}[Y]$ (inverse is not true)

•
$$var(X + Y) = var(X) + var(Y)$$

Given independent random variables X_1, \ldots, X_n ,

$$\mathit{var}(X_1 + X_2 + \cdots + X_n) = \mathit{var}(X_1) + \mathit{var}(X_2) + \cdots + \mathit{var}(X_n)$$

Some Discrete Distributions

| | Х | $p_X(k)$ | E [X] | var(X) |
|-----------|--|---|-----------------|---------------------------|
| Bernoulli | $ \begin{cases} 1 success \\ 0 failure \end{cases} $ | $\left\{ \begin{array}{ll} p & k=1\\ 1-p & k=0 \end{array} \right.$ | р | p(1-p) |
| Binomial | Number of successes in n Bernoulli trials | ${\binom{n}{k}p^k(1-p)^{n-k}}\ k=0,1,\ldots,n$ | np | np(1-p) |
| Geometric | Number of trials until first success | $(1-p)^{k-1}p\ k=1,2,\dots$ | $\frac{1}{p}$ | $\frac{1-p}{p^2}$ |
| Uniform | An integer in the interval [a,b] | $ \begin{cases} \frac{1}{b-a+1} & k = a, \dots, b \\ 0 & \text{otherwise} \end{cases} $ | <u>a+b</u> 2 | $\frac{(b-a)(b-a+2)}{12}$ |

6.041SC Probabilistic Systems Analysis and Applied Probability Fall 2013

For information about citing these materials or our Terms of Use, visit: http://ocw.mit.edu/terms.